

Math 3236 Statistical Theory

1/24/23

X_i form a random sample
with p.d.f

$$f(x; \lambda) = \lambda e^{-\lambda x}$$

$$\mathbb{E}(X_i) = \frac{1}{\lambda}$$

$$\text{Var}(X_i) = \frac{1}{\lambda^2}$$

Notation

$$\underline{X} = (X_1, \dots, X_n)$$

$$\underline{x} = (x_1, \dots, x_n)$$

Estimator for λ .

Method of moment:

$$\hat{\lambda}_n = \frac{n}{\sum_{i=1}^n X_i} = \frac{1}{\bar{X}}$$

$$\hat{\lambda}_n \xrightarrow{p} \frac{1}{\mathbb{E}(X_i)} = \lambda$$

Bias

$$S = \sum_{i=1}^N X_i$$

$$f_S(s) = \frac{\lambda^N s^{N-1} e^{-\lambda s}}{(N-1)!}$$

(Gamma dist. par. N, λ)

$$E(\hat{\lambda}_N) = \frac{N}{N-1} \lambda$$

$$\hat{\lambda}_N = \frac{N-1}{N} \hat{\lambda}_N = \frac{N-1}{\sum_i X_i} = \frac{N-1}{S}$$

$\hat{\lambda}_N$ is unbiased

$$\hat{\lambda}_N \xrightarrow{P} \lambda$$

$$E(1/S^2) = \frac{\lambda^2}{(N-1)(N-2)}$$

$$\text{Var}(\hat{\lambda}_N) = \frac{\lambda^2}{N-2}$$

You call an estimator The

MVUE

Minimum variance

unbiased estimator.

I) I call $T_N = \frac{1}{S_N}$

$$f_{T_N}(t) = \frac{\lambda^N t^{-(N+1)} e^{-\lambda/t}}{(N-1)!}$$

$$f_{\frac{1}{T_N}}(e) = \lambda^N \frac{N^{(N+1)} e^{-(N+1)} e^{-\lambda N/e}}{N!}$$

————— 0 —————
CLT

$$\lambda \sqrt{N} \left(\bar{X} - \frac{1}{\lambda} \right) \Rightarrow N(0, 1)$$

using $\varphi(x) = \frac{1}{x}$

$$\frac{\lambda \sqrt{N}}{\lambda^2} \left(\frac{1}{\bar{X}} - \lambda \right) \Rightarrow N(0, 1)$$

$$\hat{\lambda} \approx N\left(\lambda, \frac{\lambda^2}{N}\right)$$

$$\hat{\lambda}_n(\underline{X})$$

\underline{X} is the random sample

\underline{x} is a realization.

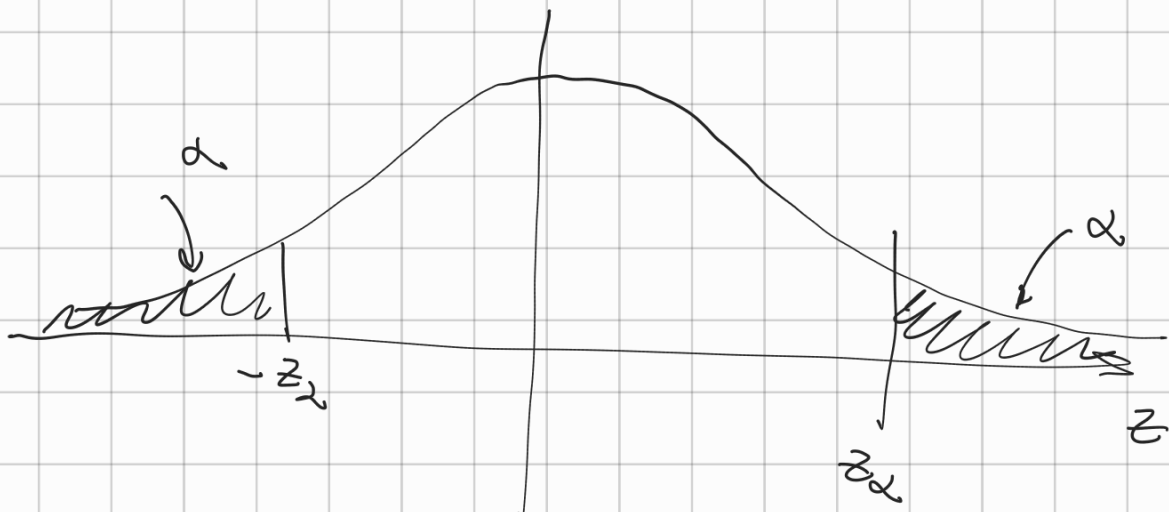
$\hat{\lambda}_n(\underline{X})$ is a r.v.
estimator

$\hat{\lambda}_n(\underline{x})$ estimate.

$$Z = \frac{\hat{\lambda}_n(\underline{X}) - \lambda}{\lambda/\sqrt{N}} \approx N(0, 1)$$

z_α a critical value

$$P(Z > z_\alpha) = \alpha$$



$$\Phi(-z_{\alpha/2}) = \alpha$$

$$P(-z_{\alpha/2} \leq Z \leq z_{\alpha/2}) = (1 - \alpha)$$

$$P\left(\hat{\lambda}(X) - z_{\alpha/2} \frac{\lambda}{\sqrt{N}} \leq \lambda \leq \hat{\lambda}(X) + \frac{\lambda z_{\alpha/2}}{\sqrt{N}}\right) = (1 - \alpha)$$

C. I. with confidence $(1 - \alpha) 100\%$.

Gamma dist.

$$\Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-\alpha x} dx$$

Prop.:

$$\Gamma(\alpha) = (\alpha-1) \Gamma(\alpha-1)$$

$$\Gamma(n) = (n-1)! \quad n \text{ integer}$$

$$\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

$$\lim_{x \rightarrow \infty} \frac{(2\pi)^{1/2} x^{x-1/2} e^{-x}}{\Gamma(x)} = 1$$

Stirling formula.

The p.d.f. of a $\Gamma(\alpha, \beta)$ r.v

$$f(x; \alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}$$

$$E(X) = \frac{\alpha}{\beta} \quad V(X) = \frac{\alpha}{\beta^2}$$

$$\psi(t) = E(e^{tx}) =$$

$$= \frac{\beta^\alpha}{\Gamma(\alpha)} \int_0^\infty e^{tx} x^{\alpha-1} e^{-\beta x} dx =$$

$$= \frac{\beta^\alpha}{\Gamma(\alpha)} \int_0^\infty x^{\alpha-1} e^{-(\beta-t)x} dx =$$

$$= \frac{\beta^\alpha}{\Gamma(\alpha)} \frac{\Gamma(\alpha)}{(\beta-t)^\alpha} = \left(\frac{\beta}{\beta-t} \right)^\alpha$$

$$E(X^k) = \left. \frac{d^k}{dt^k} \psi(t) \right|_{t=0}$$

If $X_1 \sim \Gamma(\alpha_1, \beta)$ and
 $X_2 \sim \Gamma(\alpha_2, \beta)$ incl.

$X_1 + X_2$

$$\psi_{X_1}(t) = \left(\frac{\beta}{\beta-t} \right)^{\alpha_1}$$

$$\psi_{X_2}(t) = \left(\frac{\beta}{\beta-t} \right)^{\alpha_2}$$

$$\begin{aligned} \psi_{X_1 + X_2}(t) &= \psi_{X_1}(t) \psi_{X_2}(t) = \\ &= \left(\frac{\beta}{\beta - t} \right)^{\alpha_1 + \alpha_2} \end{aligned}$$

$$X_1 + X_2 \stackrel{d}{=} \Gamma(\alpha_1 + \alpha_2, \beta)$$

if X is exp with par λ

$$X \stackrel{d}{=} \Gamma(z, \lambda)$$

$$f_X(x) = \lambda e^{-\lambda x} = \frac{\lambda^z x^{z-1} e^{-\lambda x}}{\Gamma(z)}$$

X_1 and X_2 are exp with par λ

$$X_1 + X_2 \stackrel{d}{=} \Gamma(2, \lambda)$$

$$\sum_{i=1}^n X_i \stackrel{d}{=} \Gamma(n, \lambda)$$

Beta distribution and

Bayes updates.

Coin flip.

p is unknown.

$X_1 \dots X_n$ is a random sample

$$\bar{X} = \frac{1}{n} \sum_i X_i$$

$$E(X_i) = p \implies \hat{p} = \bar{X}$$

$$E(\bar{X}) = p$$

Bayesian

Prior knowledge on p .

$$f(p) = 1 \quad 0 \leq p \leq 1$$

Now I flip my coin and I

see $H = 1$

$$f(p|z) \quad ??$$

$$f(0|z) = 0$$

$$f(p|z) = \frac{IP(z|p) f(p)}{IP(z)}$$

$$= \frac{IP(z|p) f(p)}{\int_0^1 IP(z|q) f(q) dq}$$

$$IP(z) = \int_0^1 IP(z|p) f(p) dp$$

$$IP(1|p) = p$$

$$f(p|z) = \frac{IP(1|p) f(p)}{IP(1)}$$

$$= \frac{p \cdot 1}{\int_0^1 p dp} = 2p$$

Posterior is

$$f(p|x) = zp$$

Observation

$f(p|x)$ is a p.d.f.

$P(1)$ is a number

$$f(p|0) = z(1-p)$$

$$f(p|0,1) = \frac{p(1-p)}{\int_0^1 p(1-p) dp} \rightarrow ??$$

$$= \frac{IP(0,1|p) f(p)}{IP(0,1)}$$